The Scientific Working Group on DNA Analysis Methods, better known by its acronym of SWGDAM, is a group of approximately 50 scientists representing Federal, State, and Local forensic DNA laboratories in the United States and Canada. During meetings, which are held twice a year, Committees discuss topics of interest to the forensic DNA community and often develop documents to provide direction and guidance for the community. This document was presented to SWGDAM and received approval on January 9, 2014.

The guidelines described herein supersede the Scientific Working Group on DNA Analysis Methods (SWGDAM) Y-chromosome Short Tandem Repeat (STR) Interpretation Guidelines issued in 2009. In compiling this document, a number of revisions were made to the 2009
Guidelines which are intended to make it reflective of the current state of forensic Y-STR typing. This document is not intended to invalidate any forensic Y-STR casework testing performed in accordance with the previous guidelines nor be applied retroactively. SWGDAM’s inclusion of this “retroactive” statement is intended to convey that the revised guidance be applied prospectively and not retroactively. With the underlying assumption that work (validation, training, analysis, interpretation) performed prior to the issuance of the revisions was appropriate and scientifically valid, revision of the applicable guidelines is not intended to invalidate or call into question the previous work.

Guidance is provided for forensic casework analyses on the identification and application of thresholds for allele detection and interpretation, and appropriate statistical approaches to the interpretation of Y-STR haplotypes including guidance on mixture interpretation. Laboratories are encouraged to review their standard operating procedures and validation data in light of these guidelines and to update their procedures as needed. This document is not intended to address the interpretation of analytical results from Y-STR testing using enhanced low template DNA techniques [i.e., such techniques are based on a laboratory’s internal validation and entail conditions used in the analysis of low template DNA samples that result in an increased potential for stochastic effects (such as allelic drop in or drop out, increased stutter and increased intralocus peak imbalance)].

Introduction

The interpretation of DNA typing results, to include the results of Y-STR testing, for human identification purposes requires professional judgment and expertise. Additionally, laboratories that analyze DNA samples for forensic casework purposes are required by the Quality Assurance Standards for Forensic DNA Testing Laboratories to establish and follow documented procedures for the interpretation of DNA typing results and reporting. Due to the multiplicity of forensic sample types and the potential complexity of DNA typing results, it is impractical and infeasible to cover every aspect of DNA interpretation by a preset rule. However, the laboratory should utilize written procedures for interpretation of analytical results with the understanding that specificity in the standard operating protocols will enable greater consistency and accuracy among analysts within a laboratory. It is recommended that standard operating procedures for
the interpretation of DNA typing results be sufficiently detailed that other forensic DNA analysts can review, understand in full, and assess the laboratory’s policies and practices. The laboratory’s interpretation guidelines should be based upon validation studies, scientific literature, and experience.

**Background**

Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general background information regarding forensic DNA analysis and definition of terms.

For the purposes of forensic nuclear DNA testing, the typing of autosomal STR loci should generally be performed due to their high power of discrimination and their utility for searching against the National DNA Index System. Y-STR typing is an additional tool that can be used in concert with autosomal typing for the recovery of male DNA profiles in male:female mixtures. Y-STR typing is also used in lieu of autosomal typing for the detection of male DNA in mixtures that contain an overabundance of female DNA. Given that under certain conditions a male minor contributor in a mixture of female:male DNA may only be detectable by Y-STR typing, laboratories should pursue Y-STR analysis as the most appropriate means of detecting a male contributor(s) in some forensic samples.

Due to the transmission of the Y-chromosome within a paternal lineage, Y-STR typing can also aid in the identification of missing persons.

Subsequent to a match between two samples using Y-STR testing, the Y-STR haplotype is searched against a database of Y-STR haplotypes to estimate the population frequency of the profile. It is noted that two specimens that exhibit the same Y-STR haplotype may have originated from either a common individual source, from any male within the same paternal lineage or unrelated individuals. A paternal lineage consists of those male relatives to whom the same Y-chromosome has been transmitted from a common ancestor. Barring mutation, all male relatives within the same paternal lineage have the same Y-STR profile. Attribution of the Y-STR typing results to a single individual, to the exclusion of relatives in the paternal lineage, is
generally not possible based on Y-chromosome loci. However, loci with higher mutation rates may enhance the ability to distinguish relatives in the same paternal lineage (Ballantyne et al., 2012). In addition, unrelated individuals may exhibit the same Y-STR haplotype.

1. Application of Y-STR Typing

When the detection of male DNA is relevant in the context of a case, Y-STR typing may be appropriate under certain conditions (e.g., in a mixture of male and female DNA, the alleles of male origin are shared or otherwise indistinguishable from those of female origin, or a relative abundance of female DNA precludes the detection of male alleles). Quantitative PCR that enables the estimation of the quantities of both male and female DNA is an important tool for the determination of an amplification strategy that maximizes the potential for determining the male component(s) of a mixture.

1.1 The laboratory should establish guidelines that define the parameters under which samples are subjected to Y-STR typing.

1.1.1 The ratio of the total (or female) DNA detected in a sample, relative to male DNA, can generally be predictive of the ability to detect a male minor contributor in a mixture of female:male DNA using autosomal typing.

1.1.1.1 Based on the autosomal amplification kit and detection instrumentation used, the laboratory should establish guidelines (e.g., based on total:male DNA quantities) for when detection of a male minor contributor to a mixture is not expected with autosomal typing.

1.1.2 If a ratio of total (or female) to male DNA is not used to determine the suitability of autosomal and Y-STR typing, the laboratory should establish alternative criteria for establishing amplification strategies to maximize the potential for detecting male DNA.

1.1.3 Samples for which detection of a male minor contributor is not expected with autosomal typing should be conserved for Y-STR typing
2. Preliminary Evaluation of Data

Please refer to the *SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories* for general guidance regarding the following topics: analytical threshold, internal standards, allelic ladders, controls and concordance of redundant loci.

3. Allele Designation

Alleles should be designated in accordance with recommendations of the DNA Commission of the International Society of Forensic Genetics [Gill et al. (2001), Gusmão et al. (2006)]. Please refer to the *SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories* for general guidance regarding locus and allele designation.

3.1 More than one allele can occur at any Y-STR locus in a single-source sample.

3.1.1 A multi-copy locus (e.g., DYS385 a/b) is present more than once in the genome. Accordingly, a single primer pair that amplifies a multi-copy sequence generates more than one allele. For multi-copy loci, if alleles cannot be assigned unequivocally to a particular locus (i.e., DYS385 a or DYS385 b), then both alleles should be reported (e.g., DYS385 a/b 11,14). The order that the alleles are listed (typically from the smallest to the largest) does not imply assignment to one or other of the two alternative chromosomal locations.

3.1.2 The majority of duplications (83% of those in the U.S. Y-STR Database, Release 3.0) result in alleles that differ in size by 1 repeat unit, while 17% of duplications involve 2-, 3-, 4- and partial-repeat unit differences.

3.1.3 The proximity of certain loci on the Y-chromosome allows for the simultaneous duplication of alleles at multiple loci. As an example, multi-locus duplications at the closely spaced loci DYS437, DYS439 and/or DYS389I/II are reported
within haplotypes in the U.S. Y-STR Database, as are multi-locus duplications at the closely spaced DYS439 and DYS635.

3.1.4 Tri-alleles have been reported [e.g., in the U.S. Y-STR Database, Release 3.0 (23,419 haplotypes), at DYS385 a/b (24 occurrences) and DYS390 (1 occurrence)].

3.2 Occasionally, deletion of a portion of the Y-chromosome or a primer-binding site mutation can result in the failure to detect one or more Y-STR loci. The majority of null alleles occur at a single locus per haplotype. However, multi-locus null alleles can occur due to deletion of loci that are within close proximity to one another (e.g., DYS437, DYS439 and DYS389I/II; and DYS391 and DYS635).

3.2.1 The laboratory should establish guidelines for the identification of such null alleles. As an example, this guideline may be based on experimental studies designed to distinguish a null allele from an undetected allele resulting from low template amounts, DNA degradation or inhibition.

3.2.2 A designated null allele(s) can be used for searching the US Y-STR database.

3.3 Partial Y-STR haplotypes may be used for inclusionary and exclusionary purposes. The laboratory should establish the minimum number of loci from an evidentiary profile required to perform a comparison to a reference profile.

4. Identification of Non-Allelic Peaks

Y-STR typing results generated with the current Y-STR typing kits exhibit the same non-allelic peaks observed in autosomal STR typing results. Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general guidance regarding non-allelic peaks and off-scale data.
5. Application of the Stochastic Thresholds to Allelic Peaks

Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general guidance regarding the establishment and usage of the stochastic threshold.

5.1 The stochastic threshold is not applicable to single-copy Y-STR loci.

5.1.1 Low template amplifications could have drop-out of a duplicated allele at a Y-STR locus that is typically single-copy. While duplications have been observed in single-copy Y-STR loci, they are rare. In view of this, a reasonable profile frequency estimate will generally be obtained by searching the database using the observed allele. Accounting for an undetected second allele would not result in a practical difference in probability.

5.2 The laboratory should establish a stochastic threshold for known multi-copy Y-STR loci based on empirical data derived within the laboratory and specific to the quantitation and amplification systems (e.g., kits) and the detection instrumentation used. It is noted that a stochastic threshold may be established by assessing peak height ratios across any multi-copy locus in a dilution series of DNA amplified in replicate. The RFU value above which it is reasonable to assume that, at a given locus, allelic dropout of a sister allele has not occurred constitutes a stochastic threshold.

6. Peak Height Ratio

Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general guidance regarding the establishment and usage of peak height ratio.

7. Number of Contributors to a DNA Profile

Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general guidance regarding the recognition of
mixtures, the minimum number of contributors to a mixture, and the generation of composite profiles

7.1 A specimen is generally considered to have originated from more than one male individual if two or more alleles are present at two or more loci single-copy loci.

7.2 For a given locus, assessment of repeat-unit differences in multiple alleles may aid in distinguishing a mixed sample from a single-source sample that exhibits duplication (see 3.1.1).

8. Interpretation of DNA Typing Results for Mixed Samples

Please refer to the SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories for general guidance regarding the following topics: mixture deconvolution, distinguishable and indistinguishable mixtures, mixtures with known contributors, deduced single source specimens derived from mixed specimens, the application of peak height ratio guidelines to mixture results, allele sharing, interpretation of potential stutter peaks in mixed samples and mixture ratios.

8.1 Single-source Y-STR haplotypes that are deduced from mixtures may be used for comparison purposes. Such profiles include (a) those of major (and potentially minor) contributor(s) to a distinguishable mixture, and (b) for an intimate sample, those foreign alleles derived from separation of a conditional known sample (e.g., from the victim) type.

8.2 Indistinguishable Y-STR mixtures (i.e., single-donor major and/or minor contributor haplotypes cannot be discerned) may be used for exclusionary purposes. SWGDAM has not yet reached consensus, however, on the appropriate statistical approach for estimating the occurrence of a combination of haplotypes in a population. A laboratory choosing to report inclusionary Y-STR typing results from indistinguishable mixtures that are determined to be relevant in the context of a case must perform statistical analysis in support of any inclusion. The statistical method employed must be
supported by empirical data and internal validation and documented in the laboratory’s standard operating procedures.

9. Comparison of DNA Typing Results

Please refer to the *SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories* for general guidance regarding the following topics: interpretation of evidentiary profiles relative to that of known profiles, partial profiles, possible conclusions, full accounting of mixed results, documentation of assumptions, and results for which no comparisons will be made.

10. Statistical Analysis of DNA Typing Results

Please refer to the *SWGDAM Interpretation Guidelines for Autosomal STR Typing by Forensic DNA Testing Laboratories* for general guidance regarding the following topics: when to perform statistical analysis, data appropriate for use in statistical analysis, statistical analysis for major contributors to distinguishable mixtures, reporting of statistical analysis, and documenting the source of the population database(s) used in statistical analysis.

All Y-STR loci analyzed in commercial kits are physically linked on the Y-chromosome. Due to the lack of recombination, the entire Y-chromosome haplotype must be treated as a single locus. Haplotype frequencies are estimated using the counting method. The counting method involves searching a given haplotype against a database to determine the number of times the haplotype was observed in that database. The frequency of the haplotype in the database is then estimated by dividing the count by the number of haplotypes searched.

A Y-STR profile probability can be estimated from the observed haplotype frequency by attaching a confidence interval (generally 95% or greater) to the haplotype frequency estimate to capture the effect of database size. The sampling variance of the profile probability decreases as the database size increases.
The profile probability is not the same as the match probability, which addresses the question of a match between the evidentiary and reference samples given that the reference donor is not the source of the evidentiary sample. Match probabilities are the probabilities of observing a profile given that it has already been observed and depend on the evolutionary history of the population. There are different approaches for calculating match probabilities, including coalescent theory and the theory used for autosomal profiles by Balding and Nichols (1994). These SWGDAM guidelines give a population genetic approach to calculating match probabilities, making use of the population structure parameter $\theta$. In the context of a given case, relevant population(s) for which the match probability will be estimated should be identified. A consolidated U.S. Y-STR database (http://usystrdatabase.org) has been established and may be used. Should a specific population(s) other than those provided in the U.S. Y-STR Database be required, the Y-Chromosome Haplotype Reference Database (YHRD) may be used (http://yhrd.org). A number of other Y-STR haplotype databases exist online [listed at http://www.cstl.nist.gov/biotech/strbase/y_strs.htm].

10.1 The laboratory should establish guidelines for the number of Y-STR loci used for searches of population databases. In general, the statistics for the typing results that provide the most genetic information and/or the highest discrimination potential are reported.

10.1.1 Population databases may contain profiles typed with different multiplexes, containing different numbers and/or sets of loci, such that only a subset of the database may have been typed at all of the loci present in the evidence profile. This can lead to a seemingly paradoxical situation: a more discriminating set of loci gives a less discriminating frequency estimate than a less discriminating set of loci, simply because there were fewer database profiles typed with the more discriminating set. In such situations, it is acceptable to perform additional searches of the population database using reduced locus sets in an attempt to obtain the maximal discrimination potential for that combination of evidence and population database profiles.
10.1.1.1 Regardless of the number or selection of loci searched, the most discriminating search is generally the one which gave the lowest proportion of matching haplotypes per number of profiles compared.

For example, U.S. Y-STR Database (Release 3.1) contains 1,984 samples with 23-locus profiles. Searching the database with a 23-locus evidence profile would limit comparisons to just those 1,984 profiles. Searching that same evidence profile a second time using the 11-locus subset common to all profiles in the U.S. Y-STR Database would increase the number of comparisons to 25,787 profiles.

10.1.1.2 When performing reduced locus-count searches, caution must be taken to exclude any “matches” that would have been non-matches had more of the evidence profile been searched. For example, a “match” at the 11 SWGDAM loci between the evidence and a population database sample would not be included as a match for statistical purposes if the profiles differed at any additional loci for which they both had information.

10.2 The basis for the profile probability estimation is the counting method. The application of a confidence interval accounts for database size and sampling variation.

10.2.1 The Y haplotype frequency \( (p) \) is calculated using the \( p = x/N \) formula, where \( x \) is equal to the number of times the haplotype is observed in a database containing \( N \) number of haplotypes. For example, if a haplotype has been observed twice in a database of \( N = 2000 \), the frequency of that haplotype will be: \( 2/2000 = 0.001 \).

10.2.2 Reporting a Y haplotype frequency, without a confidence interval, is acceptable but only provides a factual statement regarding observations of a Y haplotype in the database.
10.2.3 An upper confidence limit for the probability of the Y haplotype in the
population should be calculated using the method described by Clopper and
Pearson (1934). This uses the binomial distribution for the probabilities of
counts, including zero or other small numbers that are found for Y haplotypes.
If the database has \( n \) haplotypes and \( x \) of the haplotype of interest are found,
then the required upper confidence limit \( p_0 \) is the solution to

\[
\sum_{k=0}^{x} \binom{n}{k} p_0^k (1 - p_0)^{n-k} = \alpha
\]

Here \( \alpha \) gives the level of confidence: \( \alpha = 0.05 \) gives a 95\% confidence limit.
The equation finds the value \( p_0 \) of the population proportion \( p \) for which the
cumulative probability 0, 1, \ldots, \( x \) copies of the profile is equal to \( \alpha \). This
equation will require a computer to solve. A special case of the result in
equation 1 is when the haplotype of interest is not seen in the database, and \( x = 0 \). The equation now has only one term in the sum on the left hand side: \((1 - p_0)^n = \alpha\). The solution is

\[
p_0 = 1 - \alpha^{1/n}
\]

and for a 95\% confidence limit this is very close to \( 3/n \). (If \( n = 2000 \) the exact
95\% upper confidence limit is 0.0014967, whereas 3/2000 is 0.0015.)

10.2.4 Typical Clopper and Pearson (1934) upper confidence interval \( p_0 \) values at \( \alpha = 0.05 \) for various \( n \) and \( x \) are provided below for demonstrative purposes:

<table>
<thead>
<tr>
<th>( x )</th>
<th>( n = 500 )</th>
<th>( n = 1,000 )</th>
<th>( n = 5,000 )</th>
<th>( n = 10,000 )</th>
<th>( n = 50,000 )</th>
<th>( n = 100,000 )</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>0.005974</td>
<td>0.002991</td>
<td>0.000599</td>
<td>0.000300</td>
<td>0.000060</td>
<td>0.000030</td>
</tr>
<tr>
<td>1</td>
<td>0.009452</td>
<td>0.004735</td>
<td>0.000948</td>
<td>0.000474</td>
<td>0.000095</td>
<td>0.000047</td>
</tr>
<tr>
<td>2</td>
<td>0.012538</td>
<td>0.006282</td>
<td>0.001259</td>
<td>0.000629</td>
<td>0.000126</td>
<td>0.000063</td>
</tr>
<tr>
<td>3</td>
<td>0.015434</td>
<td>0.007735</td>
<td>0.001550</td>
<td>0.000775</td>
<td>0.000155</td>
<td>0.000078</td>
</tr>
<tr>
<td>4</td>
<td>0.018213</td>
<td>0.009130</td>
<td>0.001830</td>
<td>0.000915</td>
<td>0.000183</td>
<td>0.000092</td>
</tr>
<tr>
<td>5</td>
<td>0.020910</td>
<td>0.010484</td>
<td>0.002101</td>
<td>0.001051</td>
<td>0.000210</td>
<td>0.000105</td>
</tr>
</tbody>
</table>
10.3 It is recognized that population substructure exists for Y-STR haplotypes. Studies with current population databases have shown that multi-locus $\theta$ values are very small for most populations, with the magnitude of the value being inversely proportional to the number of Y-STR loci. Theta ($\theta$) is used in the following equation for the match probability:

\[
\text{Eq. 3} \quad Pr(A \mid A) = \theta + (1 - \theta) p_A
\]

where $A$ is the haplotype of interest and $Pr(A \mid A)$ is the probability of observing haplotype $A$ given that it has already been seen once in another individual of the same subpopulation. $p_A$ is the profile probability which can be estimated by the counting method, with sampling uncertainty being accommodated by using the upper confidence limit for the estimate of $p_A$.

10.3.1 Equation 3 is a match probability. It is the haplotype analog of the formulae described in National Research Council (1996) Recommendation 4.2. The match probability, or a likelihood ratio that incorporates the match probability, should be reported.

10.3.1.1 For example, a match probability of 0.001 means that there is a 1 in 1000 chance of randomly selecting a second individual with this profile given that it has already been observed once.

10.3.1.2 When the evidence profile is determined to be single source, the inverse of the match probability (Eq. 3) is the likelihood ratio. For example, a match probability of 0.001 means that the DNA match is 1000 times more likely to occur if the reference individual (or a patrilineal relative) is the contributor than if the source of the evidence is a randomly selected individual from the same population.

10.3.2 Due to an increase in the overall haplotype mutation rate, the magnitude of $\theta$ decreases as the number of loci increases. Similarly, for multiplexes with
identical numbers of loci, the magnitude of theta will be lower for the multiplex with a higher overall mutation rate.

10.3.3 Estimates of $\theta$ for Equation 3 are provided in Appendix 1. Other estimates of $\theta$ have been published [Budowle et al. (2009a), Budowle et al. (2009b), Cockerton et al. (2012), Ge et al. (2010a)].

10.3.3.1 For partial Y-STR profiles, a $\theta$ value from Appendix 1 should be selected based on the number of loci used.

10.3.4 The advent of new multiplexes (i.e., with additional Y-STR loci, or with a different set or subset of loci) will necessitate additional studies of $\theta$ which will be made available through SWGDAM.

10.3.5 Empirical studies of $\theta$ reflect the database used. $\theta$ estimates will change as the number of profiles in the database change. The changes in $\theta$ estimates are expected to be small for small proportional changes in the number of profiles.

10.4 While SWGDAM recommends the usage of match probabilities as formulated in Equation 3, alternate statistical approaches for inclusionary Y-STR typing results have been described [e.g., Brenner (2014)]. The statistical method employed by the laboratory must be supported by empirical data and internal validation and documented in the laboratory’s standard operating procedures.

10.5 The match probabilities for autosomal and Y-STR typing results obtained for a given sample may be combined. The independent inheritance of these systems suggests independence of the profile probabilities. However, there are dependencies between match probabilities among autosomal and Y profiles. These dependencies are low for populations with low ethnic heterogeneity but they increase as population structure increases. An approach to accommodate autosomal and Y profile dependencies was given by Walsh et al. (2008) and developed further by work currently in progress [preliminary results were reported by Weir et al. (2013).] Such dependencies are
negligible for match probabilities of the values currently seen when the focus is on matches to unrelated individuals, or to those more distantly related than first cousins.

10.5.1 Combining autosomal and Y-STR results requires a consistent approach to the statistical question being addressed. Under the assumption of negligible dependencies, the match probability formula (Eq. 3) is to be multiplied by the autosomal match probability as defined by National Research Council (1996) Recommendation 4.2.

10.6 Prior to reporting combined statistics for mtDNA and Y-STR results, the laboratory issuing the report should verify that each population used demonstrates independence between the mtDNA and Y-STR results. If independence cannot be demonstrated between the mtDNA and Y-STR results for the referenced population(s), combining these systems is not recommended. Ge et al. (2010b) and Roby et al. (2009) report evidence of statistical independence between mtDNA and Y-STR profiles in U.S. and Chilean populations.

10.7 The CODIS software generates a combined likelihood ratio for autosomal, mtDNA and Y-STR results for missing person searches to rank potential candidates. However, this value is not intended to be a reporting statistic.

11. References


Clopper, C. and Pearson, E. *The use of confidence or fiducial limits illustrated in the case of binomial*, Biometrika (1934) 26:404-413.


Ge, J., Chakraborty, R., and Budowle, B. *Test of independence in contingency tables of large dimension with ordered categories and its application in population genetics*, Abstract Volume of the American Society of Human Genetics 60th Annual Meeting, Washington DC (November 2-6, 2010b; Abstract # 3045/F).


Appendix 1

Estimation of $\theta$ (or $F_{ST}$) using African American, Asian, Caucasian, Hispanic, and Native American population data from the U.S. Y-STR Database release 3.2 follows from methods described by Weir and Cockerham (1984), recently modified following a suggestion of Bhatia et al. (2013). $\theta$ describes the chance of haplotypes being the same within populations relative to the chance of them being the same between populations. The original work of Weir and Cockerham assumed the same value of $\theta$ applied to all populations, and haplotype frequencies were weighted by sample sizes in calculating averages. When $\theta$ differs among populations it is more appropriate to use unweighted averages, and the analyses can be cast in terms of the proportions of pairs of haplotypes that match, within and between populations. This method was described by Weir et al. (2013), and a peer-reviewed publication is forthcoming.

To provide $\theta$ values (Table 1) for partial as well as full profiles, $\theta$ was estimated for all possible subsets of loci for each of three multiplexes (Figures 1 and 2). For convenience, $\theta$ was assumed to be the same for all haplotypes with the same number of loci. Although there is variability among single loci, the variability in $\theta$ among sets of larger numbers of loci is small, and the numerical values shown are good indicators of the effects of changing the number of loci.
Table 1. $\theta$ estimates\(^1\) for all possible subsets of loci\(^2\) for each of three multiplexes

<table>
<thead>
<tr>
<th>Loci(^3)</th>
<th>For African Americans, Asians, Caucasians &amp; Hispanics</th>
<th>For African Americans, Asians, Caucasians, Hispanics &amp; Native Americans</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0.06</td>
<td>0.06</td>
</tr>
<tr>
<td>2</td>
<td>0.04</td>
<td>0.04</td>
</tr>
<tr>
<td>3</td>
<td>0.03</td>
<td>0.03</td>
</tr>
<tr>
<td>4</td>
<td>0.02</td>
<td>0.02</td>
</tr>
<tr>
<td>5</td>
<td>0.008</td>
<td>0.008</td>
</tr>
<tr>
<td>6</td>
<td>0.005</td>
<td>0.005</td>
</tr>
<tr>
<td>7</td>
<td>0.003</td>
<td>0.003</td>
</tr>
<tr>
<td>8</td>
<td>0.002</td>
<td>0.002</td>
</tr>
<tr>
<td>9</td>
<td>0.001</td>
<td>0.002</td>
</tr>
<tr>
<td>10</td>
<td>0.0006</td>
<td>0.002</td>
</tr>
<tr>
<td>11</td>
<td>0.0004</td>
<td>0.0009</td>
</tr>
<tr>
<td>12</td>
<td>0.0002</td>
<td>0.0007</td>
</tr>
<tr>
<td>13</td>
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<td>0.0006</td>
</tr>
<tr>
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<tr>
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</tr>
<tr>
<td>22</td>
<td>0.00002</td>
<td>0.0003</td>
</tr>
</tbody>
</table>

\(^1\) Theta estimates are based on the maximum value observed among the three multiplexes shown in Figures 1 and 2.

\(^2\) Theta estimates were calculated using average matching proportions across all possible combinations of loci within each multiplex for the numbers of loci represented above (e.g., the average matching proportion for all 170,544 15-locus combinations that can occur with the PowerPlex Y23 multiplex).

\(^3\) DYS385a,b was counted as a single locus.
Figure 1. PowerPlex Y, Yfiler and PowerPlex Y23 $\theta$ estimates excluding Native American population data

Figure 2. PowerPlex Y, Yfiler and PowerPlex Y23 $\theta$ estimates including Native American population data